Time Series and Complex Networks

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Complex networks from multivariate time series



Complex networks from multivariate time series



Outline

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1 Dependence measures in univariate time series

- **O** Dependence measures in univariate time series
- **2** Interdependence in multivariate time series

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- **O** Dependence measures in univariate time series
- **2** Interdependence in multivariate time series
- **3** Complex networks from multivariate time series

- **O** Dependence measures in univariate time series
- **2** Interdependence in multivariate time series
- **3** Complex networks from multivariate time series
- **4** High-dimensional time series: Implications and solutions













A single observable Observation of an underlying system variable, $y_t = h(s_t)$			Υ _{t-5}	Υ _{t-4}	Y _{t-3}	Υ _{t-2}	Y _{t-1}	Y _t	<i>Y_{t+1}</i>
Underlying dynamical system $s_{t+1} = f(s_t)$ or $\dot{s} = f(s)$, $s_t \in \mathbb{R}^d$ or stochastic process $s_{t+1} = f(s_t, \varepsilon_t)$ or $\dot{s} = f(s, \varepsilon)$, $s_t \in \mathbb{R}^d$	1								
Reconstruction of the dynamics $y_{t+1} = F(y_t) = F(y_t, y_{t-1},, y_{t-m+1})$ or $y_{t+1} = F(y_t) = F(y_t, y_{t-\tau},, y_{t-(m-1)\tau})$ <i>m</i> : embedding dimension τ : delay time	y _t ∈ R dela	ıy en	nbeddii	ng (uni	form)				
s_t : state variable, e.g. represent y_t : observed dynamics, e.g. EEG	ing the I	brair	n activi	ţ					
	_		t-5	t-4	t-3	t-2	t-1	t	t+1

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Dependence measures in univariate time series

- Interdependence in multivariate time series
- **③** Complex networks from multivariate time series
- 4 High-dimensional time series: Implications and solutions

 $y_t = s_t$ $s_t = 1 - 1.4s_{t-1}^2 + 0.3s_{t-2}$







Are X_t and X_{t-1} linearly correlated? Are X_t and X_{t-2} linearly correlated?



3

Autocorrelation $r(\tau) = r(X_t; X_{t-\tau})$

Are X_t and X_{t-1} linearly correlated? $r(1) = r(X_t; X_{t-1}) \neq 0$? Yes Are X_t and X_{t-2} linearly correlated? $r(2) = r(X_t; X_{t-2}) \neq 0$? Yes

$$r(\tau) = r(X_t; X_{t-\tau}) = \frac{1}{n-\tau} \sum_{t=\tau+1}^n (x_t - \bar{x})(x_{t-\tau} - \bar{x})/s_X^2$$



Are X_t and X_{t-2} directly linearly correlated?



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Partial autocorrelation $\phi_{\tau,\tau} = r(X_t; X_{t-\tau} | X_{t-1}, \dots, X_{t-\tau+1})$ Are X_t and X_{t-2} directly linearly correlated? Are X_t and X_{t-2} linearly correlated given X_{t-1} ? $r(X_t; X_{t-2} | X_{t-1}) \neq 0$? No



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Are X_t and X_{t-1} linearly and nonlinearly correlated? Are X_t and X_{t-2} linearly and nonlinearly correlated?



3

Mutual information $I(\tau) = I(X_t; X_{t-\tau})$

Are X_t and X_{t-1} linearly and nonlinearly correlated? $I(X_t; X_{t-1}) \neq 0$? Yes Are X_t and X_{t-2} linearly and nonlinearly correlated? $I(X_t; X_{t-2}) \neq 0$? Yes $I(\tau) = I(X_t, X_{t-\tau}) = H(X_t) + H(X_{t-\tau}) - H(X_t, X_{t-\tau})$ $= \sum_{x_t, x_{t-\tau}} p_{X_t X_{t-\tau}}(x_t, x_{t-\tau}) \log \frac{p_{X_t X_{t-\tau}}(x_t, x_{t-\tau})}{p_{X_t}(x_t)p_{X_{t-\tau}}(x_{t-\tau})}$

$$H(X) = -\sum_{x} p(x) \log p(x)$$



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Are X_t and X_{t-2} directly linearly and nonlinearly correlated?



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Conditional mutual information $I_c(\tau) = I(X_t; X_{t-\tau}|X_{t-1}, \dots, X_{t-\tau+1})$ Are X_t and X_{t-2} directly linearly and nonlinearly correlated? Are X_t and X_{t-2} linearly and nonlinearly correlated given X_{t-1} ? $I(X_t; X_{t-2}|X_{t-1}) \neq 0$? No



(4) (3) (4) (3) (4)

Are X_t and X_{t-2} directly linearly correlated? $r(X_t; X_{t-2}|X_{t-1}) \neq 0$? No Are X_t and X_{t-2} directly linearly or/and nonlinearly correlated? $I(X_t; X_{t-2}|X_{t-1}) \neq 0$? No



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- Interdependence in multivariate time series
- Omplex networks from multivariate time series
- 4 High-dimensional time series: Implications and solutions



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Cross-correlation $r_{XY}(\tau) = r(X_t; Y_{t+\tau})$

Are X_t and Y_{t+1} linearly correlated? $r(1) = r(X_t; Y_{t+1}) \neq 0$? Yes Are X_{t-3} and Y_{t+1} linearly correlated? $r(4) = r(X_t; Y_{t+4}) \neq 0$? Yes

$$r_{XY}(\tau) = r(X_t; Y_{t+\tau}) = \frac{1}{n-\tau} \sum_{t=1}^{n-\tau} (x_t - \bar{x})(y_{t+\tau} - \bar{y})/(s_X s_Y)$$



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Cross-mutual information $I_{XY}(\tau) = I(X_t; Y_{t+\tau})$ Are X_t and Y_{t+1} linearly and nonlinearly correlated? Yes Are X_{t-3} and Y_{t+1} linearly and nonlinearly correlated? Yes

$$I_{XY}(\tau) = I(X_t, Y_{t+\tau}) = H(X_t) + H(Y_{t+\tau}) - H(X_t, Y_{t+\tau}) \\ = \sum_{x_t, y_{t+\tau}} p_{X_t Y_{t+\tau}}(x_t, y_{t+\tau}) \log \frac{p_{X_t Y_{t+\tau}}(x_t, y_{t+\tau})}{p_{X_t}(x_t) p_{Y_{t+\tau}}(y_{t+\tau})}$$



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Partial Cross-correlation $r_{XY|Z}(0) = r(X_t; Y_t|Z_t)$



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Partial Cross-correlation $r_{XY|Z}(0) = r(X_t; Y_t|Z_t)$ lags?



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Partial Cross-correlation $r_{XY|Z}(0) = r(X_t; Y_t|Z_t)$ lags? Partial Cross-mutual information $I_{XY|Z}(0) = I(X_t; Y_t|Z_t)$ lags?





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 $r_{XY}(\tau) \neq 0$ or $l_{XY}(\tau) \neq 0$: \implies correlation of x_t and $y_{t+\tau}$

 \implies X effects the future of Y, $X \rightarrow Y$

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 $r_{XY}(\tau) \neq 0$ or $I_{XY}(\tau) \neq 0$: \implies correlation of x_t and $y_{t+\tau}$ $\implies X$ effects the future of Y, $X \rightarrow Y$

$$r_{XY}(- au)
eq 0 ext{ or } I_{XY}(- au)
eq 0: \implies Y ext{ } X$$

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 $r_{XY}(\tau) \neq 0$ or $l_{XY}(\tau) \neq 0$: \implies correlation of x_t and $y_{t+\tau}$ $\implies X$ effects the future of Y, $X \rightarrow Y$

$$r_{XY}(- au)
eq 0 ext{ or } I_{XY}(- au)
eq 0: \implies Y \to X$$

Thus $r_{XY}(\tau)$ and $I_{XY}(\tau)$ indicate the direction of interaction.

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$$r_{XY}(-\tau) \neq 0 \text{ or } I_{XY}(-\tau) \neq 0: \implies Y \to X$$

Thus $r_{XY}(\tau)$ and $I_{XY}(\tau)$ indicate the direction of interaction.

Can they also be used as causality measures?

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$$r_{XY}(-\tau) \neq 0 \text{ or } I_{XY}(-\tau) \neq 0: \implies Y \to X$$

Thus $r_{XY}(\tau)$ and $I_{XY}(\tau)$ indicate the direction of interaction.

Can they also be used as causality measures? Not the most appropriate, but they have been used in many studies

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Significance randomization test for a correlation / causality measure q, H₀: q = 0 H₁: $q \neq 0$

• Generate *M* resampled (surrogate) time series, each by shifting the original observations with a random time step *w*:

original time series: $\{x_t\} = \{x_1, x_2, \dots, x_n\}$

i-th surrogate time series:

$$\{x_t^{*i}\} = \{x_{w+1}, x_{w+2}, \dots, x_n, x_1, \dots, x_{w-1}, x_w\}$$

Compute the statistic q on the original pair, q₀, and on the M surrogate pairs, q₁,..., q_M,

e.g.
$$q_0 \equiv r_{XY}(\tau) = \text{Corr}(x_t, y_{t+\tau})$$
 and $q_i \equiv \text{Corr}(x_t^{*i}, y_{t+\tau}^{*i})$

3 If q_0 is at the tails of the empirical null distribution formed by q_1, \ldots, q_M , reject H₀.

Using rank ordering: for a two-sided test, the *p*-value of the test is



Example: Returns for USA, UnitedKingdom, Greece and Australia. Correlation matrix for delay 1, $r_{XY}(1)$

$$R(1) = \begin{bmatrix} 0.382 & 0.333 & 0.596 \\ 0.049 & 0.039 & 0.303 \\ 0.096 & 0.001 & 0.190 \\ 0.031 & -0.001 & -0.021 \end{bmatrix}$$

Randomization significance test for $r_{XY}(1)$ (M = 1000) Matrix of *p*-values Adjacency

Adjacency matrix

$$P(R(1)) = \begin{bmatrix} 0.0013 & 0.0013 & 0.0033\\ 0.0732 & 0.1991 & 0.0013\\ 0.0073 & 0.8901 & 0.0033\\ 0.2450 & 0.9760 & 0.4028 \end{bmatrix} A = \begin{bmatrix} 1 & 1 & 1 \\ 0 & 0 & 1 \\ 1 & 0 & 1 \\ 0 & 0 & 0 \end{bmatrix}$$

For significance level, say $\alpha = 0.05$, there may be $p < \alpha$ more often than it should be due to multiple testing. Correction with e.g. False Discovery Rate (FDR) Example: Returns for USA, UnitedKingdom, Greece and Australia. Correlation matrix for delay 1, $r_{XY}(1)$



Randomization significance test for $r_{XY}(1)$ (M = 1000)Matrix of p-valuesAdjacency matrix

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For significance level, say $\alpha = 0.05$, there may be $p < \alpha$ more often than it should be due to multiple testing. Correction with e.g. False Discovery Rate (FDR)

Linear causality measures (direct and indirect)

Idea of Granger causality $X \rightarrow Y$ [Granger 1969]:

predict Y better when including X in the regression model.

Granger Causality Index (GCI) [Brandt & Williams 2007]

Bivariate time series $\{x_t, y_t\}_{t=1}^n$ driving system: X, response system: Y

Model 1 (restricted, R, X absent in the model):

$$y_t = \sum_{i=1}^p a_i y_{t-i} + e_{R,t}$$

Model 2 (unrestricted, U, X present in the model):

$$y_t = \sum_{i=1}^{p} a_i y_{t-i} + \sum_{i=1}^{p} b_i x_{t-i} + e_{U,t}$$

$$\operatorname{Var}(\hat{e}_{R,t}) = a_i y_{t-i} + a_{i-1} + a$$

 $\operatorname{\mathsf{GCl}}_{X\to Y} = \operatorname{\mathsf{In}} \frac{\operatorname{\mathsf{var}}(e_{R,t})}{\operatorname{\mathsf{Var}}(\hat{e}_{U,t})} \qquad \operatorname{\mathsf{GCl}}_{X\to Y} > 0 \Rightarrow X \to Y \text{ holds}$

 $\operatorname{GCl}_{X \to Y} > 0$? \Rightarrow Significance test

If X does not Granger causes Y then the contribution of X-lags in the unrestricted model should be insignificant \Rightarrow

the terms of X should be insignificant

H₀: $b_i = 0$, for all i = 1, ..., pH₁: $b_i \neq 0$, for any of i = 1, ..., p

Snedecor-Fisher test (F-test):

$$F = \frac{(SSE^R - SSE^U)/p}{SSE^U/ndf}$$

SSE: sum of squared errors ndf: number of degrees of freedoms, ndf = (n - p) - 2p, n - p: number of equations, 2p: number of coefficients in the U-model.

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Linear causality measures (direct)

Conditional Granger Causality Index (CGCI)

K time series $\{x_t, y_t\}_{t=1}^n$ and $\{\mathbf{z}_t\}_{t=1}^n = \{z_{1,t}, z_{2,t}, \dots, z_{K-2,t}\}_{t=1}^n$ driving system: *X*, response system: *Y*, conditioning on system *Z*, $Z = \{Z_1, Z_2, \dots, Z_{K-2}\}$

Model 1 (restricted, R, X absent in the model):

$$y_t = \sum_{i=1}^{p} a_i y_{t-i} + \sum_{i=1}^{p} A_i \mathbf{z}_{t-i} + e_{R,t}$$

Model 2 (unrestricted, U, X present in the model):

$$y_t = \sum_{i=1}^{p} a_i y_{t-i} + \sum_{i=1}^{p} b_i x_{t-i} + \sum_{i=1}^{p} A_i \mathbf{z}_{t-i} + e_{U,t}$$
$$\mathsf{CGCl}_{X \to Y|Z} = \mathsf{In} \frac{\mathsf{Var}(\hat{e}_{R,t})}{\mathsf{Var}(\hat{e}_{U,t})}$$

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 $CGCI_{X \to Y|Z} > 0 \quad ? \quad \Rightarrow \quad \text{Significance test as for GCI}$ $H_0: \ b_i = 0, \text{ for all } i = 1, \dots, p$ $H_1: \ b_i \neq 0, \text{ for any of } i = 1, \dots, p$ $F = \frac{(\text{SSE}^R - \text{SSE}^U)/p}{\text{SSE}^U/\text{ndf}}$

ndf = (n - p) - Kp, n - p: number of equations, Kp: number of coefficients in the U-model.

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VAR model for Y

$$y_t = \sum_{i=1}^{p} a_i y_{t-i} + \sum_{i=1}^{p} b_i x_{t-i} + e_{U,t}$$

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VAR model for Y

$$y_t = \sum_{i=1}^{p} a_i y_{t-i} + \sum_{i=1}^{p} b_i x_{t-i} + e_{U,t}$$

$$y_{t+1} = \sum_{i=1}^{p} a_i y_{t-i+1} + \sum_{i=1}^{p} b_i x_{t-i+1} + e_{U,t+1}$$

 $\begin{aligned} &y_{t+1} \text{ is given in terms of } \mathbf{y}_t = [y_t, y_{t-1}, \dots, y_{t-p+1}] \text{ and} \\ &\mathbf{x}_t = [x_t, x_{t-1}, \dots, x_{t-p+1}], \quad y_{t+1} = \mathbf{F}(\mathbf{y}_t, \mathbf{x}_t) \end{aligned}$

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VAR model for Y

$$y_t = \sum_{i=1}^{p} a_i y_{t-i} + \sum_{i=1}^{p} b_i x_{t-i} + e_{U,t}$$

$$y_{t+1} = \sum_{i=1}^{p} a_i y_{t-i+1} + \sum_{i=1}^{p} b_i x_{t-i+1} + e_{U,t+1}$$

 y_{t+1} is given in terms of $\mathbf{y}_t = [y_t, y_{t-1}, \dots, y_{t-p+1}]$ and $\mathbf{x}_t = [x_t, x_{t-1}, \dots, x_{t-p+1}], \quad y_{t+1} = \mathbf{F}(\mathbf{y}_t, \mathbf{x}_t)$ Let the lag step be $\tau \ge 1 \implies \mathbf{y}_t = [y_t, y_{t-\tau}, \dots, y_{t-(p-1)\tau}]$: τ, p : embedding parameters (generally different for X and Y)

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VAR model for Y

$$y_t = \sum_{i=1}^{p} a_i y_{t-i} + \sum_{i=1}^{p} b_i x_{t-i} + e_{U,t}$$

$$y_{t+1} = \sum_{i=1}^{p} a_i y_{t-i+1} + \sum_{i=1}^{p} b_i x_{t-i+1} + e_{U,t+1}$$

 $\begin{array}{l} y_{t+1} \text{ is given in terms of } \mathbf{y}_t = [y_t, y_{t-1}, \ldots, y_{t-p+1}] \text{ and} \\ \mathbf{x}_t = [x_t, x_{t-1}, \ldots, x_{t-p+1}], \quad y_{t+1} = \mathbf{F}(\mathbf{y}_t, \mathbf{x}_t) \\ \text{Let the lag step be } \tau \geq 1 \quad \Rightarrow \quad \mathbf{y}_t = [y_t, y_{t-\tau}, \ldots, y_{t-(p-1)\tau}]: \\ \tau, p: \text{ embedding parameters (generally different for X and Y)} \end{array}$

State space reconstruction:

 $\begin{aligned} \mathbf{x}_t &= [x_t, x_{t-\tau_x}, \dots, x_{t-(m_x-1)\tau_x}]', \text{ embedding parameters: } & m_x, \tau_x \\ \mathbf{y}_t &= [y_t, y_{t-\tau_y}, \dots, y_{t-(m_y-1)\tau_y}]', \text{ embedding parameters: } & m_y, \tau_y \end{aligned}$

 y_{t+1} : future state of Y

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Transfer Entropy (TE) [Schreiber, 2000]

Measure the effect of X on Y at one time step ahead, accounting (conditioning) for the effect from its own current state

$$\mathsf{TE}_{X \to Y} = I(y_{t+1}; \mathbf{x}_t | \mathbf{y}_t)$$

= $H(\mathbf{x}_t, \mathbf{y}_t) - H(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t) + H(y_{t+1}, \mathbf{y}_t) - H(\mathbf{y}_t)$
= $\sum p(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t) \log \frac{p(y_{t+1} | \mathbf{x}_t, \mathbf{y}_t)}{p(y_{t+1} | \mathbf{y}_t)}$

Transfer Entropy (TE) [Schreiber, 2000]

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Joint entropies (and distributions) can have high dimension! Entropy estimates from nearest neighbors [Kraskov et al, 2004]

Transfer Entropy (TE) [Schreiber, 2000]

Measure the effect of X on Y at one time step ahead, accounting (conditioning) for the effect from its own current state

$$TE_{X \to Y} = I(y_{t+1}; \mathbf{x}_t | \mathbf{y}_t)$$

= $H(\mathbf{x}_t, \mathbf{y}_t) - H(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t) + H(y_{t+1}, \mathbf{y}_t) - H(\mathbf{y}_t)$
= $\sum p(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t) \log \frac{p(y_{t+1} | \mathbf{x}_t, \mathbf{y}_t)}{p(y_{t+1} | \mathbf{y}_t)}$

Joint entropies (and distributions) can have high dimension!

Entropy estimates from nearest neighbors [Kraskov et al, 2004]

TE is equivalent to GCI when the stochastic process of (X, Y) is Gaussian [Barnett et al, PRE 2009]
driving system: X, response system: Y, conditioning on system Z, $Z = \{Z_1, Z_2, \dots, Z_{K-2}\}$ join all K - 2 z-reconstructed vectors: $\mathbf{Z}_t = [\mathbf{z}_{1,t}, \dots, \mathbf{z}_{K-2,t}]$

Measure the effect of X on Y at T times ahead, accounting (conditioning) for the effect from its own current state and the current state of the other variables except X.

Partial Transfer Entropy (PTE) [Vakorin et al, 2009; Papana et al, 2012]

$$\begin{aligned} \mathsf{PTE}_{X \to Y|Z} &= I(y_{t+1}; \mathbf{x}_t | \mathbf{y}_t, \mathbf{Z}_t) \\ &= H(\mathbf{x}_t, \mathbf{y}_t | \mathbf{Z}_t) - H(y_{t+1}, \mathbf{x}_t, \mathbf{y}_t | \mathbf{Z}_t) + H(y_{t+1}, \mathbf{y}_t | \mathbf{Z}_t) - H(\mathbf{y}_t | \mathbf{Z}_t) \end{aligned}$$

Joint entropies (and distributions) can have very high dimension!

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Example: Nonlinear stochastic process

Nonlinear stochastic map:

$$\begin{aligned} x_{1,t} &= 3.4x_{1,t-1}(1-x_{1,t-1}^2)e^{-x_{1,t-1}^2} + 0.4e_{1,t} \\ x_{2,t} &= 3.4x_{2,t-1}(1-x_{2,t-1}^2)e^{-x_{2,t-1}^2} + 0.5x_{1,t-1}x_{2,t-1} + 0.4e_{2,t} \\ x_{3,t} &= 3.4x_{3,t-1}(1-x_{3,t-1}^2)e^{-x_{3,t-1}^2} + 0.3x_{2,t-1} + 0.5x_{1,t-1}^2 + 0.4e_{3,t} \end{aligned}$$

[Model 7, Gourevich et al, 2006]

True connectivity network



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[Model 7, Gourevich et al, 2006]

True connectivity network



Estimation of the correct causality effects from the time series?

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Example: Games of world cup 1930 - 2006



Example: Flight connections



Source: https://au.pinterest.com/pin/488077678338752549/

Example: Greek domestic flight connections



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Example: Bullying causal effects



Source: https://doi.org/10.1093/schbul/sbx013

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Example: Global financial market

MSCI market capitalization weighted index



Data source: https://www.msci.com/market-cap-weighted-indexes

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Data source: https://physionet.org/pn6/chbmit/chb08/

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Time Series and Networks



Three possible ways to convert a network of weighted connections (the Granger causality measure) to a network of binary connections:



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① Threshold on the measure magnitude, $q(i \rightarrow j) > \text{thr.}$

Three possible ways to convert a network of weighted connections (the Granger causality measure) to a network of binary connections:



- **①** Threshold on the measure magnitude, $q(i \rightarrow j) > \text{thr.}$
- ② Threshold on the network density, only the d% largest q(i → j).

Three possible ways to convert a network of weighted connections (the Granger causality measure) to a network of binary connections:



- **①** Threshold on the measure magnitude, $q(i \rightarrow j) > \text{thr.}$
- ② Threshold on the network density, only the d% largest $q(i \rightarrow j)$.
- Significance test on each $q(i \rightarrow j)$. Threshold, e.g. $\alpha = 0.05$ on the *p*-value of the test.

Parametric or resampling test (resampling test for a nonlinear causality measure).

Significance resampling test on $q(i \rightarrow j)$ for each pair (X_i, X_j) .

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Popular choice:

False Discovery Rate (FDR) [Benjamini & Hochberg, 1995]

• $\mathcal{K}(\mathcal{K}-1)$ *p*-values in ascending order: $p_{(1)}, p_{(2)}, \dots, p_{(\mathcal{K}(\mathcal{K}-1))}$

Rejection for the k tests with p ≤ p_(k), where p_(k) is the largest p-value for which p_(k) < kα/(K(K − 1)).

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Small *p*-value can only be obtained with large number of surrogates When K gets large, FDR requires **huge** M (impractical).

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$$\begin{aligned} x_{1,t+1} &= 1.4 - x_{1,t}^2 + 0.3x_{1,t-1} \\ x_{i,t+1} &= 1.4 - (0.5C(x_{i-1,t} + x_{i+1,t}) + (1-C)x_{i,t})^2 + 0.3x_{i,t-1} \\ x_{K,t+1} &= 1.4 - x_{K,t}^2 + 0.3x_{K,t-1} \end{aligned}$$

C: coupling strength [Politi & Torcini, 1992]

Network structure for K = 5



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Example, TE, K = 5



Kugiumtzis Dimitris

- Dependence measures in univariate time series
- Interdependence in multivariate time series
- **③** Complex networks from multivariate time series
- I High-dimensional time series: Implications and solutions

Example, TE, K = 10

True network

Binary network from Threshold (thr=0.01)



The curse of dimensionality:

• For FDR, in general $M \sim K(K-1)/\alpha$. When K gets large, huge M may be required (impractical).

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Example, TE, K = 20

True network





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• For K > 2, bivariate measures are likely to produce false couplings (indirect connections).

• Multivariate measures require long time series, e.g. $PTE_{X \to Y|Z} = I(y_{t+1}; \mathbf{x}_t | \mathbf{y}_t, \mathbf{Z}_t)$ requires the estimation of entropy of $[y_{t+1}, \mathbf{x}_t, \mathbf{y}_t, \mathbf{Z}_t]'$ of dimension 1 + Km.

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Example, PTE, K = 4

Binary network from Threshold (thr=0.01) True network Weighted network from PTE(m=2,tau=1) Binary network from Density (dens=0.30) Binary network from Significance (alpha=0.050) Binary network from FDR-Signficance (alpha=0.050 time step t

Example, PTE, K = 8



Interdependence using Dimension Reduction

K time series $\{x_t, y_t\}_{t=1}^n$ and $\{\mathbf{z}_t\}_{t=1}^n = \{z_{1,t}, z_{2,t}, \dots, z_{K-2,t}\}_{t=1}^n$ driving system: X, response system: Y, conditioning on system Z, $Z = \{Z_1, Z_2, \dots, Z_{K-2}\}$

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If K large with respect to n multivariate Granger causality is problematic ("the curse of dimensionality").

In CGCI, the VAR model

$$y_{t} = \sum_{i=1}^{p} a_{i} y_{t-i} + \sum_{i=1}^{p} b_{i} x_{t-i} + \sum_{i=1}^{p} A_{i} \mathbf{z}_{t-i} + e_{U,t}$$

has Kp lagged variables.

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has Kp lagged variables.

Statistics methodology: dimension reduction, sparse regression, restricted regression, and sparse/retricted VAR models.

Partial Mutual Information from Mixed Embedding (PMIME) applies dimension reduction using mutual information. The idea [Vlachos & Kugiumtzis, 2010]:

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Partial Mutual Information from Mixed Embedding (PMIME) applies dimension reduction using mutual information. The idea [Vlachos & Kugiumtzis, 2010]:

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Similar approaches based on this idea: [Faes et al, PRE 2011; Stamaglia et al, PRE 2012; Runge et al, PRL 2012; Wibral et al, PLOSOne 2013; Runge et al, PRE 2015; edited book of Wibral, Vicente and Lizier "Directed information measures in Neuroscience", Springer, 2014.]

The mixed embedding scheme

Start with an empty embedding vector w⁰_t, future vector of Y, y_{t+1}, and maximum lag L (or L_x for X, L_y for Y etc)
 W_t = {x_t,..., x_{t-L-1}, y_t,..., y_{t-L-1}, z_{1,t},..., z_{K-2,t-L-1}}

- Start with an empty embedding vector \mathbf{w}_t^0 , future vector of Y, y_{t+1} , and maximum lag L (or L_x for X, L_y for Y etc) $\mathbf{W}_t = \{x_t, \dots, x_{t-L-1}, y_t, \dots, y_{t-L-1}, z_{1,t}, \dots, z_{K-2,t-L-1}\}$
- First embedding cycle: $w_t^1 = \operatorname{argmax}_{w \in W_t} I(y_{t+1}; w)$, and $\mathbf{w}_t^1 = (w_t^1)$

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- At embedding cycle j suppose $\mathbf{w}_t^{j-1} = (w_t^1, w_t^2, \dots, w_t^{j-1}).$

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- At embedding cycle j suppose $\mathbf{w}_t^{j-1} = (w_t^1, w_t^2, \dots, w_t^{j-1})$. Add $w_t^j \in \mathbf{W}_t \setminus \mathbf{w}_t^{j-1}$ that maximizes mutual information to y_{t+1} conditioning on the current \mathbf{w}_t^{j-1} ,

- Start with an empty embedding vector \mathbf{w}_t^0 , future vector of Y, y_{t+1} , and maximum lag L (or L_x for X, L_y for Y etc) $\mathbf{W}_t = \{x_t, \dots, x_{t-L-1}, y_t, \dots, y_{t-L-1}, z_{1,t}, \dots, z_{K-2,t-L-1}\}$
- First embedding cycle: $w_t^1 = \operatorname{argmax}_{w \in W_t} I(y_{t+1}; w)$, and $\mathbf{w}_t^1 = (w_t^1)$
- At embedding cycle j suppose w_t^{j-1} = (w_t¹, w_t², ..., w_t^{j-1}). Add w_t^j ∈ W_t \ w_t^{j-1} that maximizes mutual information to y_{t+1} conditioning on the current w_t^{j-1}, w_t^j = argmax_{w∈mathbfWt}\w^{j-1}/(y_{t+1}; w|w_t^{j-1})

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- First embedding cycle: $w_t^1 = \operatorname{argmax}_{w \in W_t} I(y_{t+1}; w)$, and $\mathbf{w}_t^1 = (w_t^1)$
- At embedding cycle j suppose w_t^{j-1} = (w_t¹, w_t², ..., w_t^{j-1}). Add w_t^j ∈ W_t \ w_t^{j-1} that maximizes mutual information to y_{t+1} conditioning on the current w_t^{j-1}, w_t^j = argmax<sub>w∈mathbfW_t\w_t^{j-1} I(y_{t+1}; w|w_t^{j-1})
 </sub>
- Progressive vector building stops at step j ($\mathbf{w}_t = \mathbf{w}_t^{j-1}$): Criterion of hard threshold: $I(y_{t+1}; \mathbf{w}_t^{j-1})/I(y_{t+1}; \mathbf{w}_t^j) > A$ (here A = 0.95) Criterion of adaptive threshold: randomization significance test on $I(y_{t+1}; w_t^j | \mathbf{w}_t^{j-1})$

The non-uniform mixed embedding vector of lags of all X, Y, Z for explaining y_{t+1} :

$$\mathbf{w}_{t} = \left(\underbrace{x_{t-\tau_{x1}}, \ldots, x_{t-\tau_{xm_x}}}_{\mathbf{w}_{t}^{x}}, \underbrace{y_{t-\tau_{y1}}, \ldots, y_{t-\tau_{ym_y}}}_{\mathbf{w}_{t}^{y}}, \underbrace{z_{t-\tau_{z1}}, \ldots, z_{t-\tau_{zm_z}}}_{\mathbf{w}_{t}^{z}}\right)$$

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The causality measure PMIME

$$R_{X \to Y|Z} = \frac{I(y_{t+1}; \mathbf{w}_t^X \mid \mathbf{w}_t^y, \mathbf{w}_t^Z)}{I(y_{t+1}; \mathbf{w}_t)}$$

*R*_{X→Y|Z}: information on the future of Y explained only by X-components of the embedding vector (given the components of Y and Z), normalized with the mutual information of the future of Y and the embedding vector.

The non-uniform mixed embedding vector of lags of all X, Y, Z for explaining y_{t+1} :

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The causality measure PMIME

$$R_{\boldsymbol{X} \to \boldsymbol{Y}|\boldsymbol{Z}} = \frac{I(y_{t+1}; \boldsymbol{w}_t^{\boldsymbol{X}} \mid \boldsymbol{w}_t^{\boldsymbol{Y}}, \boldsymbol{w}_t^{\boldsymbol{Z}})}{I(y_{t+1}; \boldsymbol{w}_t)}$$

- *R*_{X→Y|Z}: information on the future of *Y* explained only by *X*-components of the embedding vector (given the components of *Y* and *Z*), normalized with the mutual information of the future of *Y* and the embedding vector.
- If \mathbf{w}_t contains no components from X, then $R_{X \to Y|Z} = 0$ and X has no direct effect on the future of Y.

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R_{X→Y|Z} = 0 when no significant causality is present, and *R_{X→Y|Z}* > 0 when it is present
 [no significance test, no issues with multiple testing!]

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 \Rightarrow good candidate for causality analysis with many variables

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Coupled identical Mackey-Glass delayed differential equations

$$\dot{x}_i(t) = -0.1 x_i(t) + \sum_{j=1}^{K} rac{C_{ij} x_j(t-\Delta)}{1+x_j(t-\Delta)^{10}} \quad ext{for} \quad i=1,\ldots,K$$

K = 5



Mackey-Glass,
$$C = 0.2$$

 $\Delta = 20$



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Mackey-Glass,
$$C = 0.2$$

 $\Delta = 20$



 $\Delta = 100$



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Kugiumtzis Dimitris Time Series and Networks

Can different network structures be detected?

Simulation: three types of networks for the generating system



Can different network structures be detected?

Simulation: three types of networks for the generating system



Generating system: coupled Mackey-Glass system, K = 25, $\Delta = 100$, C = 0.2with coupling structure defined by the network type

Causality measure: PMIME

Estimation of the Random Network



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Estimation of the Small-World Network



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Estimation of the Scale-Free Network



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The network structure undergoes structural change at specific time points: Random \Rightarrow Small-World \Rightarrow Scale-Free

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Estimation of networks with PMIME at sliding windows

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Estimation of networks with PMIME at sliding windows

Estimation of network characteristics on the PMIME networks

The network structure undergoes structural change at specific time points: Random \Rightarrow Small-World \Rightarrow Scale-Free

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Structural change detection,
[Slow], [Middle], [Fast], [Very fast]
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EEG and Transcranial Magnetic Stimulation (TMS)



Jointly with Vassilis Kimiskidis, Medical School, AUTh



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TMS-EEG: brain connectivity analysis

How does TMS act on epileptic brain connectivity?

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How does TMS act on epileptic brain connectivity?

 $\{X_1, X_2, \dots, X_K\}$: *K* EEG channels, each represents a (sub)system Practical problems to overcome:

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Practical problems to overcome:

• Application on small time windows \Rightarrow limited data size

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PMIME addresses all these problems!

Example: compare PMIME to other measures on EEG

[Kugiumtzis, PRE, 2013]

One epileptiform discharge (ED) episode terminated by transcranial magnetic stimulation (TMS), totally 45 channels

- Select randomly a subset of channels.
- Compute the connectivity measures on the subset at each sliding window
- Compute average connectivity strength at each sliding window.
- Repeat the steps above a number of times (here 12).

... for subsets of 5, 15, 25, 35 and once for 45 channels.



Kugiumtzis Dimitris

Epileptiform discharges induced by TMS

Preprocessing:

[One Episode]

- replacement of TMS artifact, high order FIR
- rejection of channels with artifacts
- reference to infinity (REST) [Qin et al, ClinNeuroph 2010]
- overlapping windows of 2s, a sliding step of 0.5s



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subject 1 with focal seizure, ED episode ends with TMS

In-Strength

Out-Strength

[Kugiumtzis & Kimiskidis, 2015]



Average strength / degree



In-Out-Strength



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Subject 1 with focal seizure, 13 episodes, average degree



structure as if it would have terminated spontaneously



structure as if it would have terminated spontaneously

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Many network indices (totally 78) computed on the PMIME-causality networks

Symbol	Description
deg^m	degree distribution, m=mean,std,skewness,kurtosis
strm	strength distribution, m=mean,std,skewness,kurtosis
TrR_k	transitivity ratio, k=binary undirected (bu), binary directed (bd)
	weighted directed (wd)
EigC ^m	eigenvector centrality distribution, m=mean,std
λ_k	characteristic path length, k=bd,wd
GE_k	global efficiency, k=bd,wd
ϵ_k^m	eccentricity distribution, m=mean,std and k=bd,wd
radk	radius, k=bd,wd
d_k	diameter, k=bd,wd
C_k^m	clustering coefficient distribution,m=mean,std and k=bd,wd
g_k^m	betweenness centrality distribution,m=mean,std and k=bd,wd
$e - g_k^m$	edge betweenness centrality distribution,m=mean,std and k=bd,wd
LE_k^m	local efficiency distribution,m=mean,std and k=bd,wd
3motif(i)	ith motif of 3 nodes, i=1,2,13
modul(i)	modularity for i modules, i=2,3,5
$r_{deg}(i, j)$	assortativity coefficient in terms of the degree, i=in,out and j=in,out or i,j=und
$r_{str}(i, j)$	assortativity coefficient in terms of the strength, i=in,out and j=in,out or i,j=und
Ptop	Rent exponent: topological
p_{ph}	Rent exponent: physical
Pee	Rent exponent:efficient embedding
SW_k	small-worldness, k=bd,wd
kcs	k-core size, k=90-percentile of degree distribution
SCS	s-core size, k=90-percentile of strength distribution
ϕ_k	Rich club coefficient, k=bd,wd
cycprob ₁	fraction of 3-cycles out of 3-paths
$cycprob_2$	probability: non-cyclic 2-path extend to 3-cycle

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Many network indices (totally 78) computed on the PMIME-causality networks

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Symbol	Description		
deg ^m	degree distribution, m=mean,std,skewness,kurtosis		
str ^m	strength distribution, m=mean,std,skewness,kurtosis	E. L. L.	1.1.1.1.1.1
TrR_k	transitivity ratio, k=binary undirected (bu), binary directed (bd)	For both s	subjects
	weighted directed (wd)	and pairs	
EigC ^m	eigenvector centrality distribution, m=mean,std	anu pans	
λ_k	characteristic path length, k=bd,wd	preED - E	D
GE_k	global efficiency, k=bd,wd	p. e	-
ϵ_k^m	eccentricity distribution, m=mean,std and k=bd,wd	ED - post	ED
rad_k	radius, k=bd,wd	· · · · · · · · · · · · · · · · · · ·	
d_k	diameter, k=bd,wd	Measure	AUROC
C_k^m	clustering coefficient distribution,m=mean,std and k=bd,wd	J mean	0.0207
S_k^m	betweenness centrality distribution,m=mean,std and k=bd,wd	aeginaan	0.9296
$e - g_k^m$	edge betweenness centrality distribution,m=mean,std and k=bd,wd	d_{bd}	0.9268
LE_k^m	local efficiency distribution,m=mean,std and k=bd,wd	3motif(1)	0.0231
3motif(i)	i^{m} motif of 3 nodes, $i=1,2,13$	$Smon_{J}(1)$	0.9231
modul(i)	modularity for i modules, i=2,3,5	λ_{bd}	0.9231
$r_{deg}(i, j)$	assortativity coefficient in terms of the degree, i=in,out and j=in,out or i,j=und	strmean	0.9207
$r_{str}(l, J)$	assortativity coefficient in terms of the strength, i=in,out and j=in,out or i,j=und	$2 \dots (C(2))$	0.0207
Ptop	Rent exponent: topological	5monf(3)	0.9206
p_{ph}	Rent exponent: physical	3motif(5)	0.9199
Pee	Rent exponent:efficient embedding	I Emean	0.0173
SWk	small-worldness, k=bd,wd	LLbd	0.9173
KCS	k-core size, k=90-percentile of degree distribution	GE_{wd}	0.9163
SCS	s-core size, k=90-percentile of strength distribution	TrR	0.9150
φ_k	Kich club coefficient, K=bd, wd	1 , 1 Wa	0.7100
cycprob ₁	inaction of 5-cycles out of 5-paths		
cycprob ₂	probability: non-cyclic 2-path extend to 3-cycle		

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Many network indices (totally 78) computed on the PMIME-causality networks

Symbol	Description		
deg^m	degree distribution, m=mean,std,skewness,kurtosis		
str ^m	strength distribution, m=mean,std,skewness,kurtosis	E. L. L.	1.1.1.1.1
TrR_k	transitivity ratio, k=binary undirected (bu), binary directed (bd)	For both s	ubjects
	weighted directed (wd)	and pairs	
EigC ^m	eigenvector centrality distribution, m=mean,std	anu pans	
λ_k	characteristic path length, k=bd,wd	preED - E	D
GE_k	global efficiency, k=bd,wd		
ϵ_k^m	eccentricity distribution, m=mean,std and k=bd,wd	ED - post	ED
rad_k	radius, k=bd,wd		
d_k	diameter, k=bd,wd	Measure	AUROC
C_k^m	clustering coefficient distribution,m=mean,std and k=bd,wd	1 - mean	0.0207
g_k^m	betweenness centrality distribution,m=mean,std and k=bd,wd	aeginaan	0.9296
$e - g_k^m$	edge betweenness centrality distribution,m=mean,std and k=bd,wd	d_{bd}	0.9268
LE_k^m	local efficiency distribution,m=mean,std and k=bd,wd	3motif(1)	0.0231
3motif(i)	i^{th} motif of 3 nodes, $i=1,2,\dots 13$	Sinon f(1)	0.9251
modul(i)	modularity for i modules, i=2,3,5	Abd	0.9231
$r_{deg}(l, J)$	assortativity coefficient in terms of the degree, i=in,out and j=in,out or i,j=und	strmean	0.9207
$r_{str}(l, j)$	Pant associativity coefficient in terms of the strength, i=in,out and j=in,out or i,j=und	3motif(3)	0.0206
Ptop	Rent exponent obvical	Shidt f(3)	0.9200
P ph	Rent exponent efficient embedding	3motif(5)	0.9199
SW1	small-worldness, k=bd.wd	LEmean	0.9173
kcs	k-core size, k=90-percentile of degree distribution	GE^{ba}	0.9163
SCS	s-core size, k=90-percentile of strength distribution	O Lwd	0.9109
ϕ_k	Rich club coefficient, k=bd,wd	TrR _{wd}	0.9150
cycprob ₁	fraction of 3-cycles out of 3-paths		

cycprob2 probability: non-cyclic 2-path extend to 3-cycle

Many networks indices discriminate well preED, ED, postED

subject 1 with genetic generalized epilepsy (GGE)

ED induced by TMS, [PMIME on 2s windows]

[Kugiumtzis et al, 2016]

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subject 1 with genetic generalized epilepsy (GGE)

ED induced by TMS, **[PMIME on 2s windows]** One episode

[Kugiumtzis et al, 2016]



Kugiumtzis Dimitris

subject 1 with genetic generalized epilepsy (GGE)

ED induced by TMS, [PMIME on 2s windows] One episode

[Kugiumtzis et al, 2016]

29 episodes

Average strength



Kugiumtzis Dimitris

PMIME on 2s window

S1D2RESTPMIME, window[2sec, 0.5sec overlap]: 489/489, time: 246/246sec causality measure: PMIMEsig, network measure: mean strength



Kugiumtzis Dimitris

$\mathsf{RGPDC}(\alpha)$ on 1s window

S1D2REST, window[1sec, 0.25sec overlap]: 981/981, time: 246/246sec causality measure: RGPDCa, network measure: mean strength Key timologal whole record time[sec], zoom 12sec: cyan box mean strength 240 mean strength

Kugiumtzis Dimitris

iCoh(lpha) on 1s window



Kugiumtzis Dimitris

Synchronization Likelihood (SL) on 1s window

S1D2REST, window[1sec, 0.25sec overlap]: 981/981, time: 246/246sec causality measure: SL, network measure: mean strength



Kugiumtzis Dimitris

ED induced by TMS, SL on 1s windows,

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ED induced by TMS, SL on 1s windows,



One episode

ED induced by TMS, SL on 1s windows,



ED induced by TMS, SL on 1s windows,



ED induced by TMS, SL on 1s windows,





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• "Bivariate measures are short for forming complex networks"

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- "Multivariate measures are hard to estimate"

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- PMIME is model-free and almost parameter-free, can estimate nonlinear direct causal effects in the presence of many variables

- "Bivariate measures are short for forming complex networks"
- "Multivariate measures are hard to estimate"
- PMIME is model-free and almost parameter-free, can estimate nonlinear direct causal effects in the presence of many variables
- How can we learn the underlying dynamics of high-dimensional time series?



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